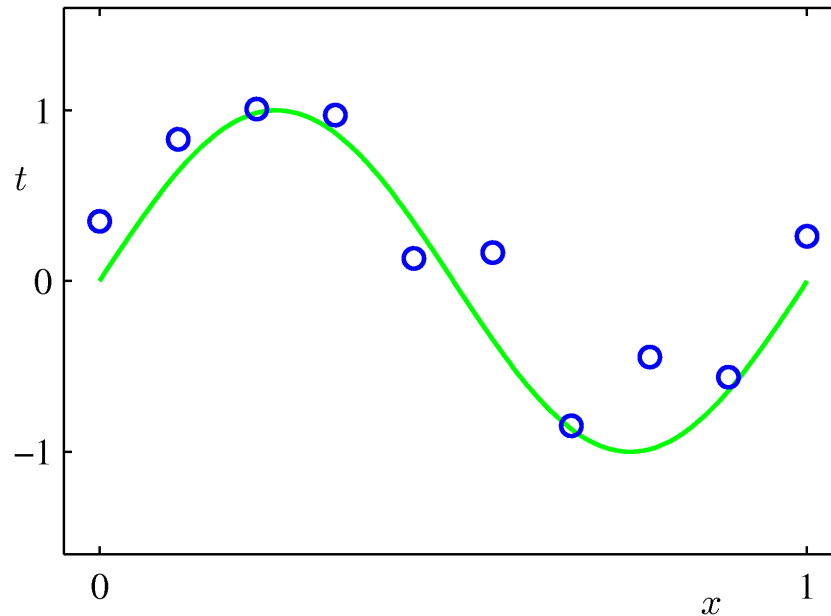


Linear Basis Function Models (1)

Example: Polynomial Curve Fitting



$$y(x, \mathbf{w}) = w_0 + w_1x + w_2x^2 + \dots + w_Mx^M = \sum_{j=0}^M w_jx^j$$

Linear Basis Function Models (2)

Generally

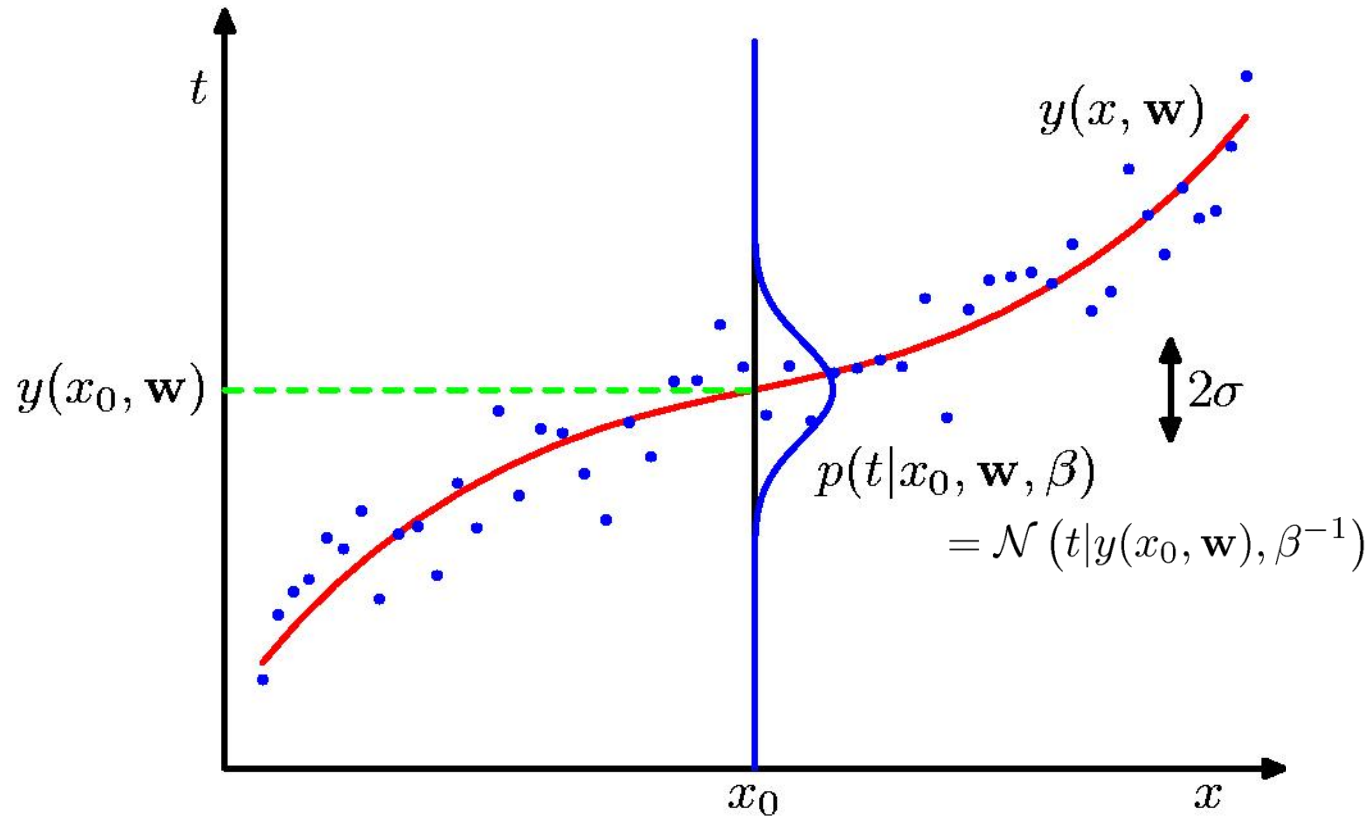
$$y(\mathbf{x}, \mathbf{w}) = \sum_{j=0}^{M-1} w_j \phi_j(\mathbf{x}) = \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x})$$

Where $\phi_j(\mathbf{x})$ are known as *basis functions*.

Typically, $\phi_0(\mathbf{x}) = 1$, so that w_0 acts as a bias.

In the simplest case, we use linear basis functions : $\phi_d(\mathbf{x}) = x_d$.

Curve Fitting Re-visited



Maximum Likelihood and Least Squares (1)

Assume observations from a deterministic function with added Gaussian noise:

$$t = y(\mathbf{x}, \mathbf{w}) + \epsilon \quad \text{where} \quad p(\epsilon|\beta) = \mathcal{N}(\epsilon|0, \beta^{-1})$$

which is the same as saying,

$$p(t|\mathbf{x}, \mathbf{w}, \beta) = \mathcal{N}(t|y(\mathbf{x}, \mathbf{w}), \beta^{-1}).$$

Given observed inputs, $\mathbf{X} = \{\mathbf{x}_1, \dots, \mathbf{x}_N\}$, and targets, $\mathbf{t} = [t_1, \dots, t_N]^T$, we obtain the likelihood function

$$p(\mathbf{t}|\mathbf{X}, \mathbf{w}, \beta) = \prod_{n=1}^N \mathcal{N}(t_n|\mathbf{w}^T \phi(\mathbf{x}_n), \beta^{-1}).$$

Maximum Likelihood and Least Squares (2)

Taking the logarithm, we get

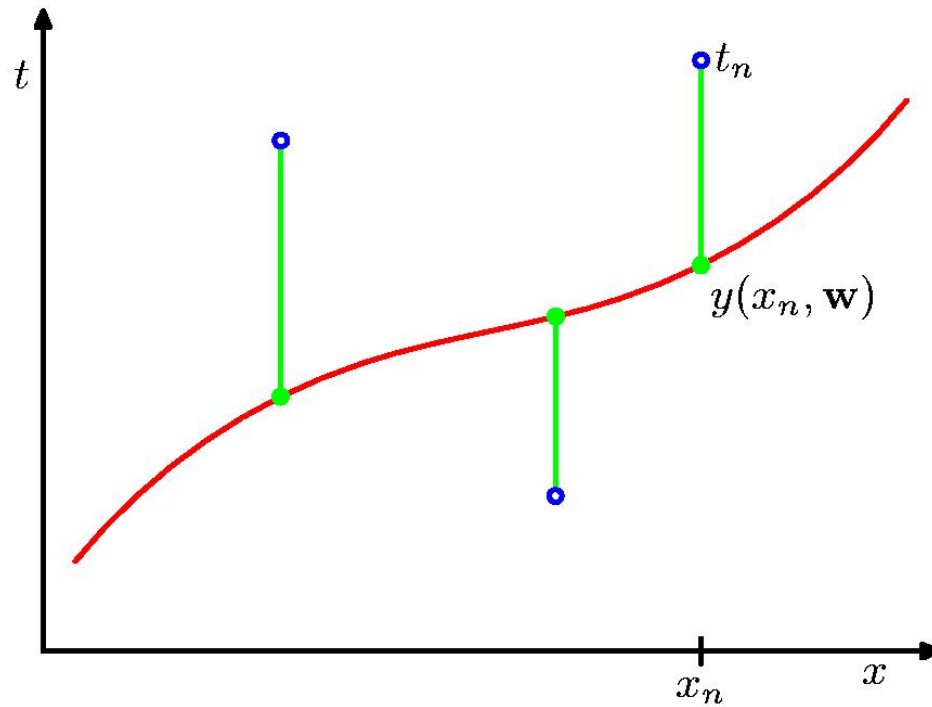
$$\begin{aligned}\ln p(\mathbf{t}|\mathbf{w}, \beta) &= \sum_{n=1}^N \ln \mathcal{N}(t_n | \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n), \beta^{-1}) \\ &= \frac{N}{2} \ln \beta - \frac{N}{2} \ln(2\pi) - \beta E_D(\mathbf{w})\end{aligned}$$

where

$$E_D(\mathbf{w}) = \frac{1}{2} \sum_{n=1}^N \{t_n - \mathbf{w}^T \boldsymbol{\phi}(\mathbf{x}_n)\}^2$$

is the sum-of-squares error.

Sum-of-Squares Error Function



Maximum Likelihood and Least Squares (3)

Computing the gradient and setting it to zero yields

$$\nabla_{\mathbf{w}} \ln p(\mathbf{t}|\mathbf{w}, \beta) = \beta \sum_{n=1}^N \{t_n - \mathbf{w}^T \phi(\mathbf{x}_n)\} \phi(\mathbf{x}_n)^T = \mathbf{0}.$$
$$0 = \sum_{n=1}^N t_n \phi(\mathbf{x}_n)^T - \mathbf{w}^T \left(\sum_{n=1}^N \phi(\mathbf{x}_n) \phi(\mathbf{x}_n)^T \right)$$

Solving for \mathbf{w} , we get

$$\mathbf{w}_{\text{ML}} = \left(\Phi^T \Phi \right)^{-1} \Phi^T \mathbf{t}$$

The Moore-Penrose pseudo-inverse, Φ^\dagger .

where

$$\Phi = \begin{pmatrix} \phi_0(\mathbf{x}_1) & \phi_1(\mathbf{x}_1) & \cdots & \phi_{M-1}(\mathbf{x}_1) \\ \phi_0(\mathbf{x}_2) & \phi_1(\mathbf{x}_2) & \cdots & \phi_{M-1}(\mathbf{x}_2) \\ \vdots & \vdots & \ddots & \vdots \\ \phi_0(\mathbf{x}_N) & \phi_1(\mathbf{x}_N) & \cdots & \phi_{M-1}(\mathbf{x}_N) \end{pmatrix}.$$

Regularized Least Squares (1)

Consider the error function:

$$E_D(\mathbf{w}) + \lambda E_W(\mathbf{w})$$

Data term + Regularization term

With the sum-of-squares error function and a quadratic regularizer, we get

$$\frac{1}{2} \sum_{n=1}^N \{t_n - \mathbf{w}^T \phi(\mathbf{x}_n)\}^2 + \frac{\lambda}{2} \mathbf{w}^T \mathbf{w}$$

which is minimized by

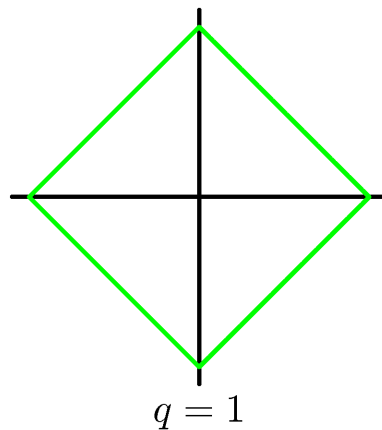
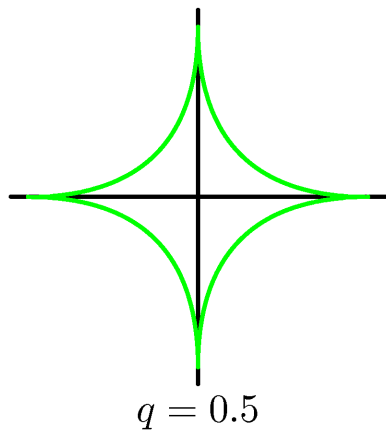
$$\mathbf{w} = \left(\lambda \mathbf{I} + \Phi^T \Phi \right)^{-1} \Phi^T \mathbf{t}.$$

λ is called the regularization coefficient.

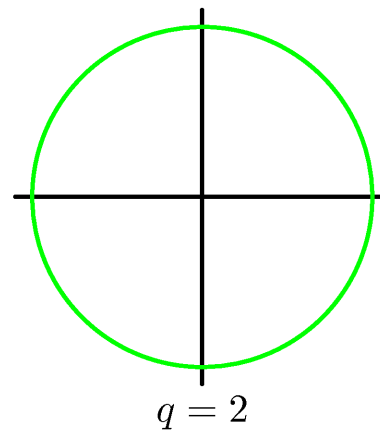
Regularized Least Squares (2)

With a more general regularizer, we have

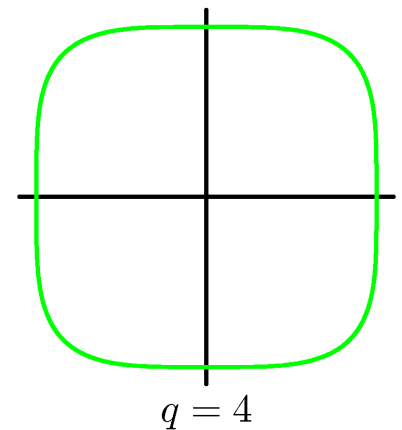
$$\frac{1}{2} \sum_{n=1}^N \{t_n - \mathbf{w}^T \phi(\mathbf{x}_n)\}^2 + \frac{\lambda}{2} \sum_{j=1}^M |w_j|^q$$



Lasso

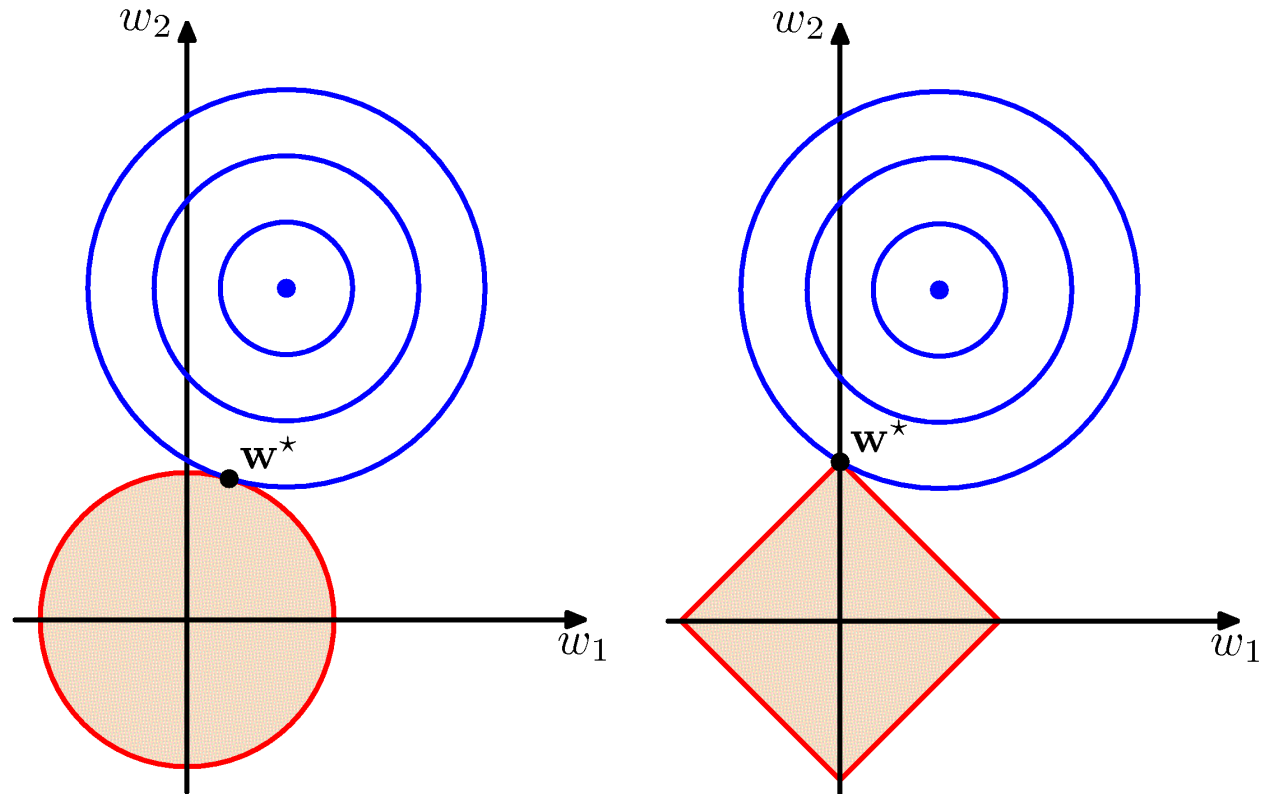


Quadratic



Regularized Least Squares (3)

Lasso tends to generate sparser solutions than a quadratic regularizer.



Bayesian Linear Regression (1)

Define a conjugate prior over \mathbf{w}

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w} | \mathbf{m}_0, \mathbf{S}_0).$$

Combining this with the likelihood function and using results for marginal and conditional Gaussian distributions, gives the posterior

$$p(\mathbf{w} | \mathbf{t}) = \mathcal{N}(\mathbf{w} | \mathbf{m}_N, \mathbf{S}_N)$$

where

$$\mathbf{m}_N = \mathbf{S}_N \left(\mathbf{S}_0^{-1} \mathbf{m}_0 + \beta \Phi^T \mathbf{t} \right)$$

$$\mathbf{S}_N^{-1} = \mathbf{S}_0^{-1} + \beta \Phi^T \Phi.$$

Bayesian Linear Regression (2)

A common choice for the prior is

$$p(\mathbf{w}) = \mathcal{N}(\mathbf{w} | \mathbf{0}, \alpha^{-1} \mathbf{I})$$

for which

$$\begin{aligned} \mathbf{m}_N &= \beta \mathbf{S}_N \Phi^T \mathbf{t} \\ \mathbf{S}_N^{-1} &= \alpha \mathbf{I} + \beta \Phi^T \Phi. \end{aligned}$$

The log of the posterior distribution is given by the sum of the log likelihood and the log of the prior

$$\ln p(\mathbf{w} | \mathbf{t}) = -\frac{\beta}{2} \sum_{n=1}^N \{t_n - \mathbf{w}^T \phi(\mathbf{x}_n)\}^2 - \frac{\alpha}{2} \mathbf{w}^T \mathbf{w} + \text{const.}$$

Bayesian Linear Regression (2)

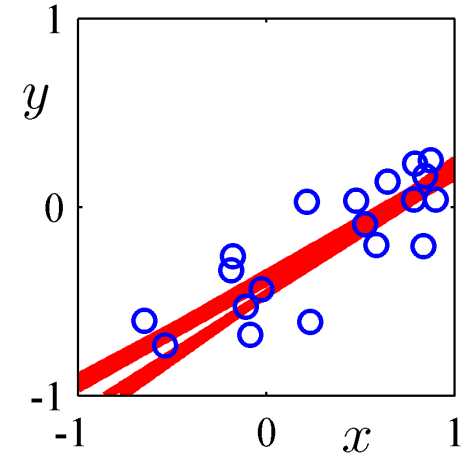
Example: estimate linear model of the form

$$y(x, \mathbf{w}) = w_0 + w_1 x$$

Data: draw x_n from uniform distribution, then plug into

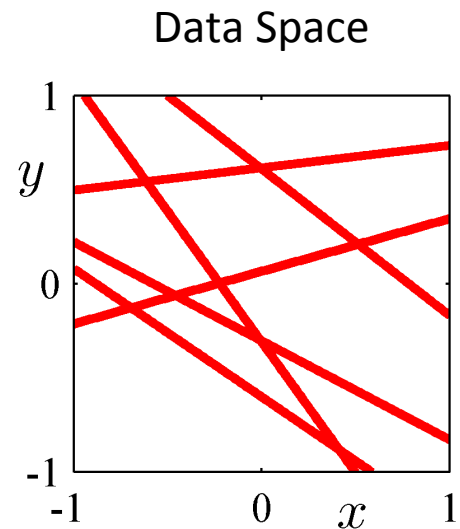
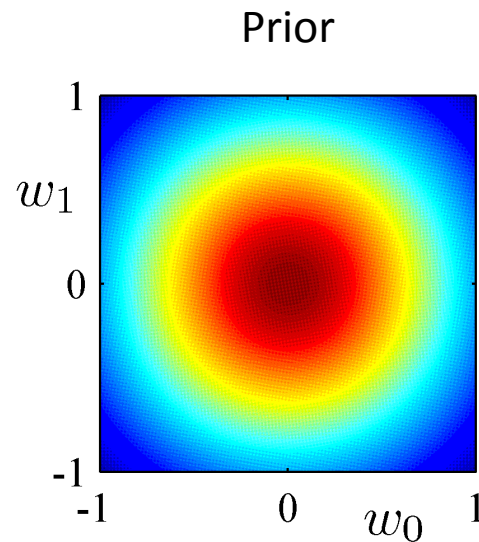
$$f(x, \mathbf{a}) = a_0 + a_1 x$$

then add Gaussian noise to obtain target value t_n



Bayesian Linear Regression (3)

0 data points observed

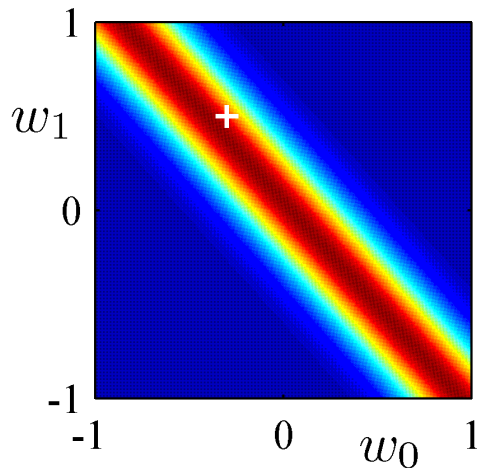


$$y(x, \mathbf{w}) = w_0 + w_1 x$$

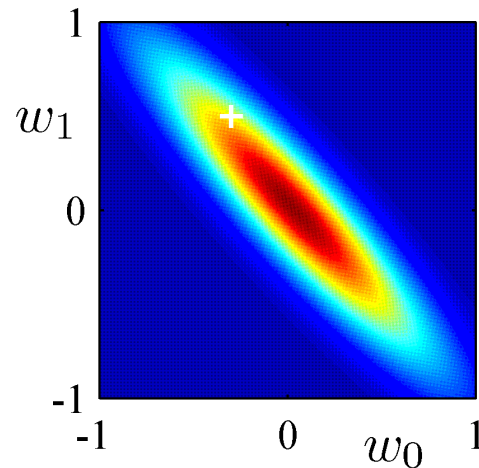
Bayesian Linear Regression (4)

1 data point observed

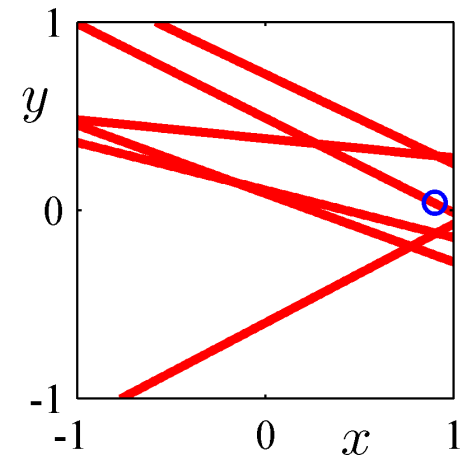
Likelihood



Posterior



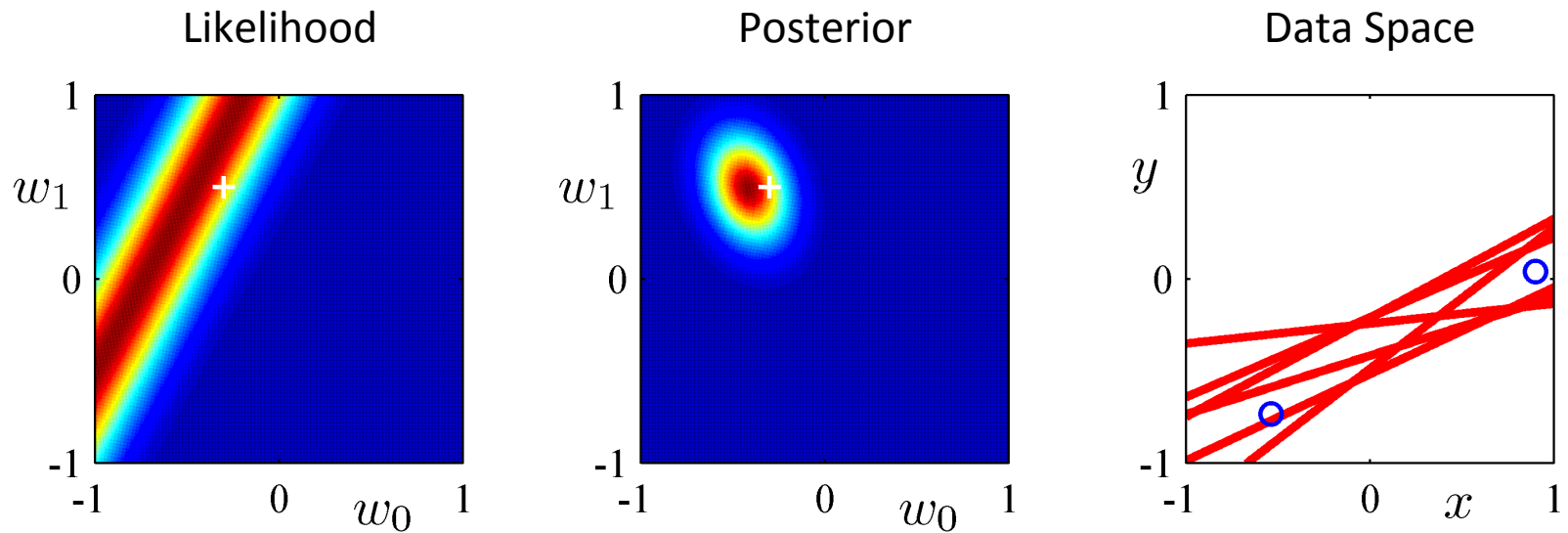
Data Space



$$y(x, \mathbf{w}) = w_0 + w_1 x$$

Bayesian Linear Regression (5)

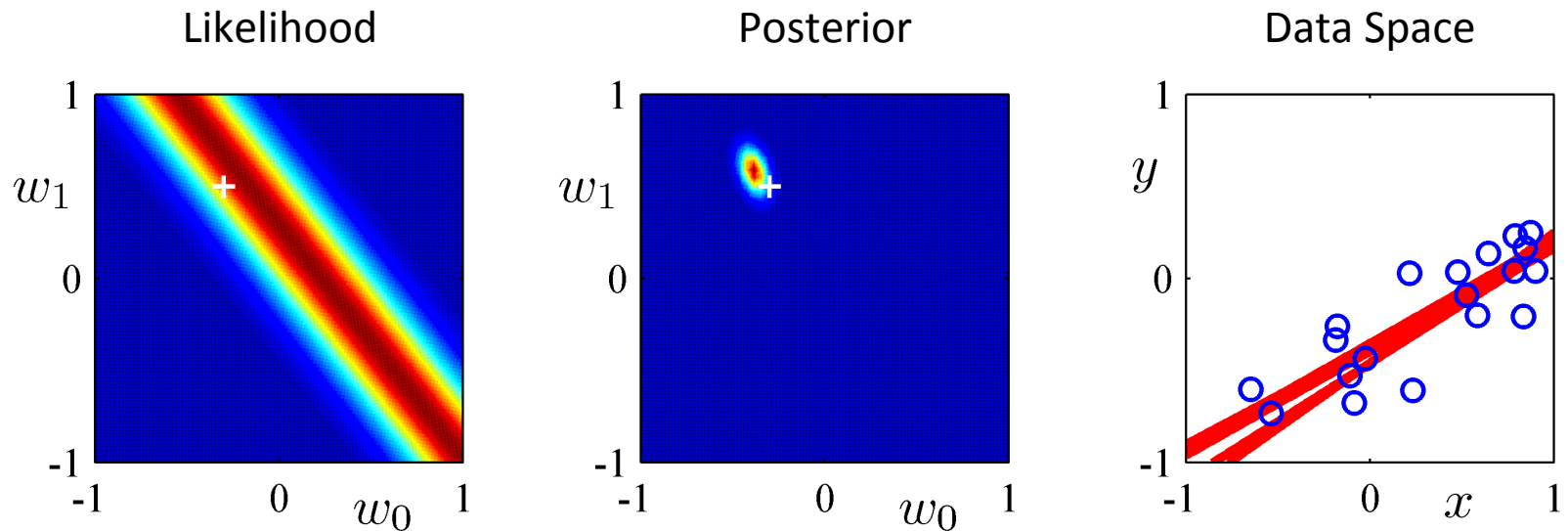
2 data points observed



$$y(x, \mathbf{w}) = w_0 + w_1 x$$

Bayesian Linear Regression (6)

20 data points observed



$$y(x, \mathbf{w}) = w_0 + w_1 x$$

Predictive Distribution (1)

Predict t for new values of \mathbf{x} by integrating over \mathbf{w} :

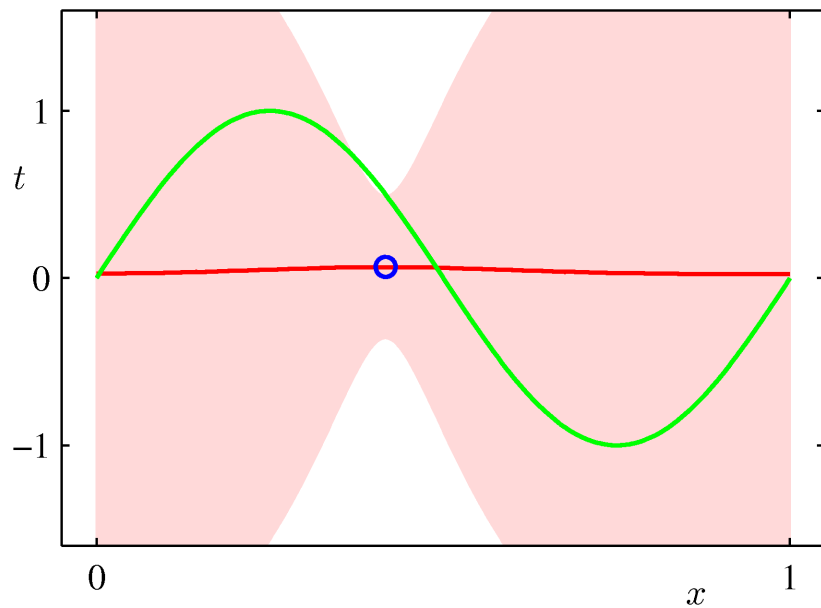
$$\begin{aligned} p(t|\mathbf{t}, \alpha, \beta) &= \int p(t|\mathbf{w}, \beta) p(\mathbf{w}|\mathbf{t}, \alpha, \beta) d\mathbf{w} \\ &= \mathcal{N}(t|\mathbf{m}_N^T \phi(\mathbf{x}), \sigma_N^2(\mathbf{x})) \end{aligned}$$

where

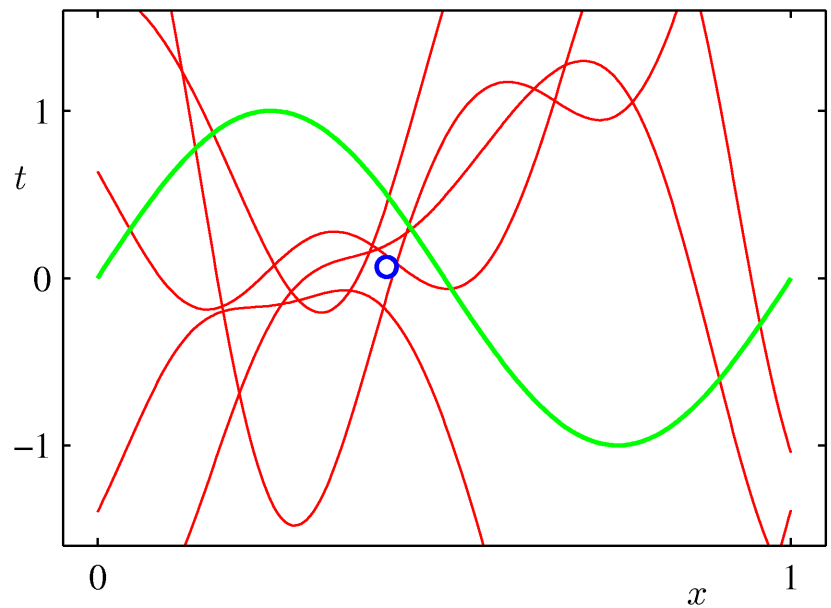
$$\sigma_N^2(\mathbf{x}) = \frac{1}{\beta} + \phi(\mathbf{x})^T \mathbf{S}_N \phi(\mathbf{x}).$$

Predictive Distribution (2)

Example: Sinusoidal data, 9 Gaussian basis functions,
1 data point



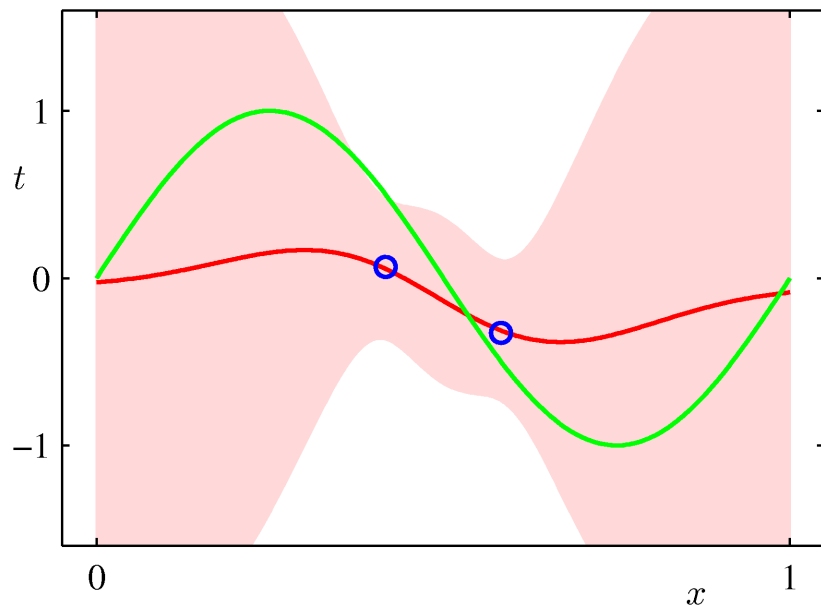
$$\mathcal{N}(t | \mathbf{m}_N^T \phi(\mathbf{x}), \sigma_N^2(\mathbf{x}))$$



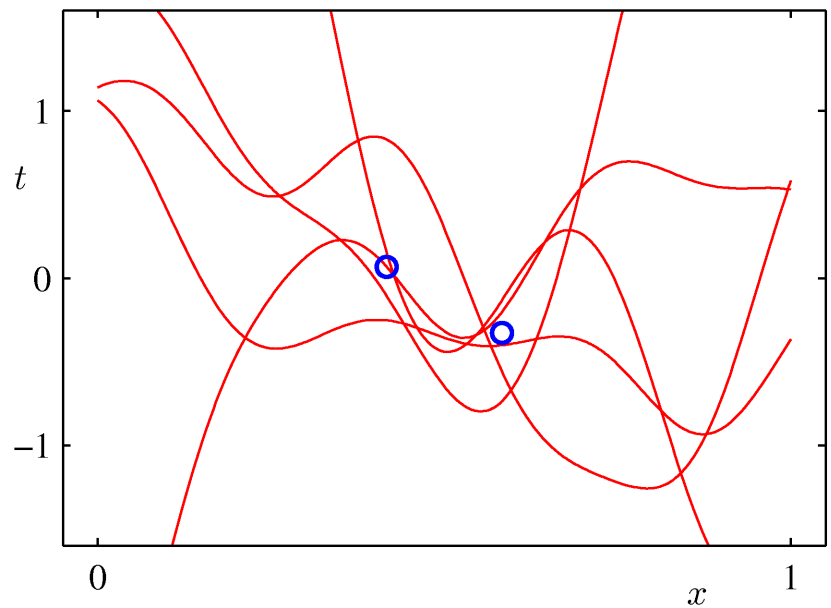
$$y(x, \mathbf{w})$$

Predictive Distribution (3)

Example: Sinusoidal data, 9 Gaussian basis functions, 2 data points



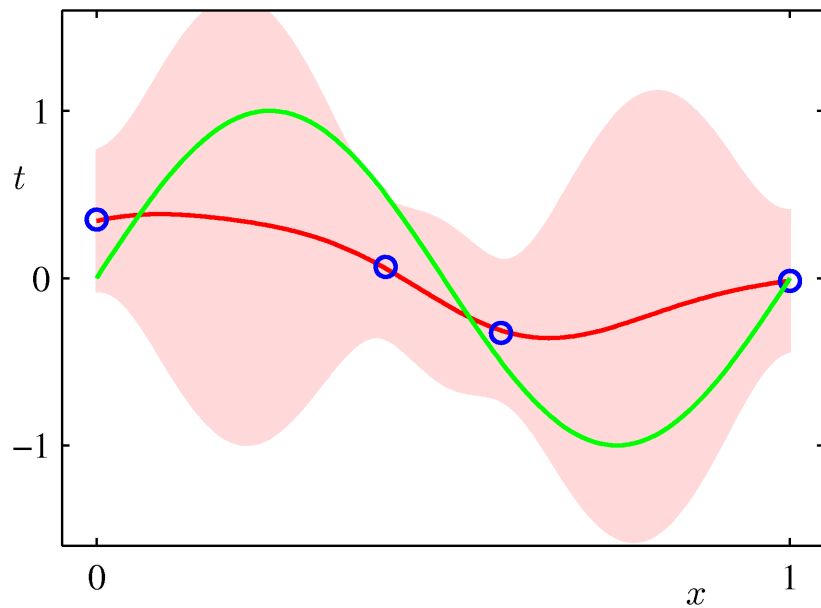
$$\mathcal{N}(t | \mathbf{m}_N^T \phi(\mathbf{x}), \sigma_N^2(\mathbf{x}))$$



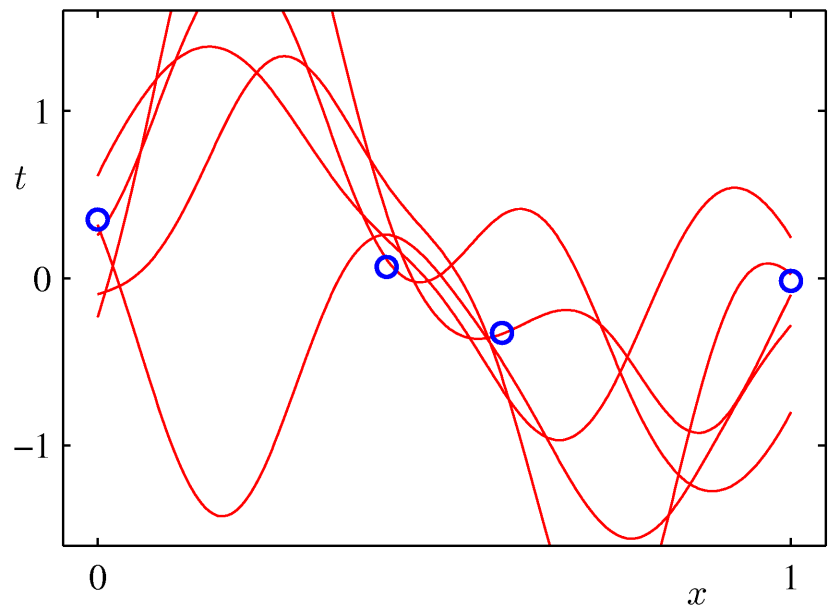
$$y(x, \mathbf{w})$$

Predictive Distribution (4)

Example: Sinusoidal data, 9 Gaussian basis functions, 4 data points



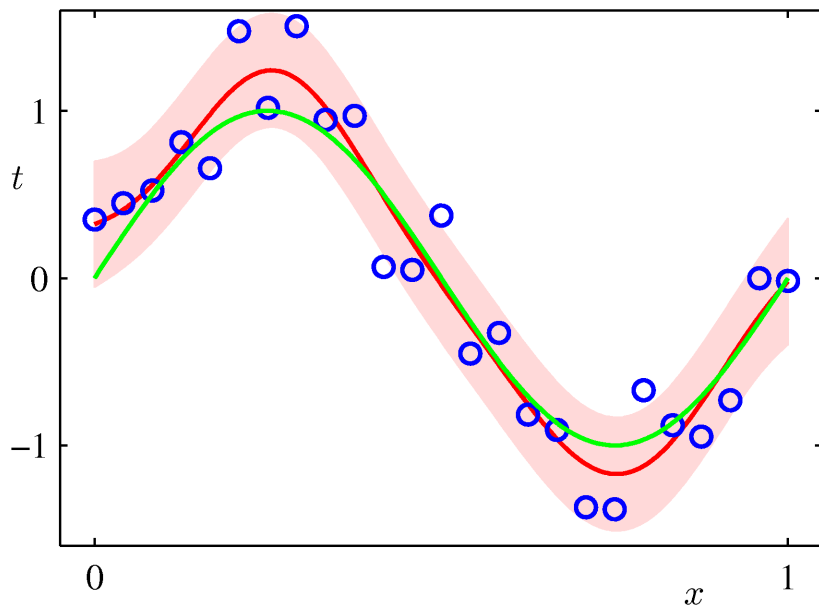
$$\mathcal{N}(t | \mathbf{m}_N^T \phi(\mathbf{x}), \sigma_N^2(\mathbf{x}))$$



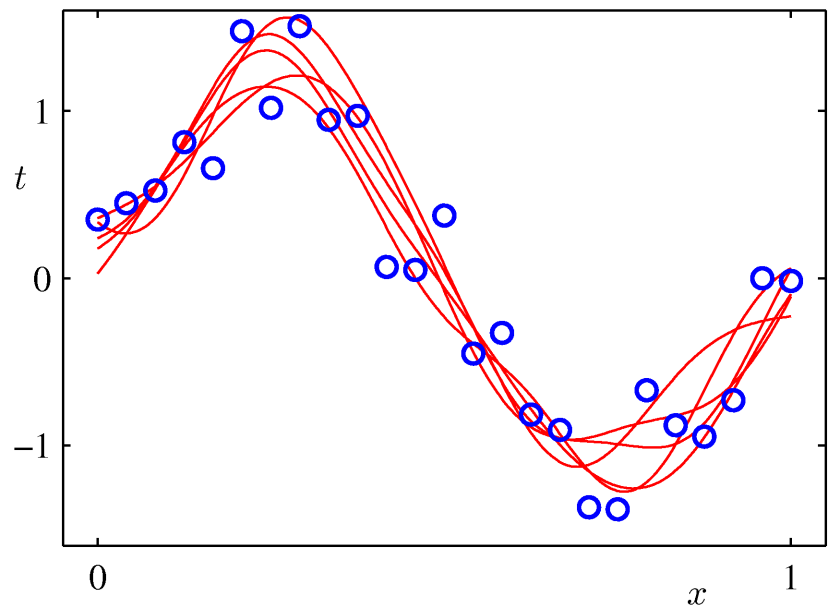
$$y(x, \mathbf{w})$$

Predictive Distribution (5)

Example: Sinusoidal data, 9 Gaussian basis functions, 25 data points



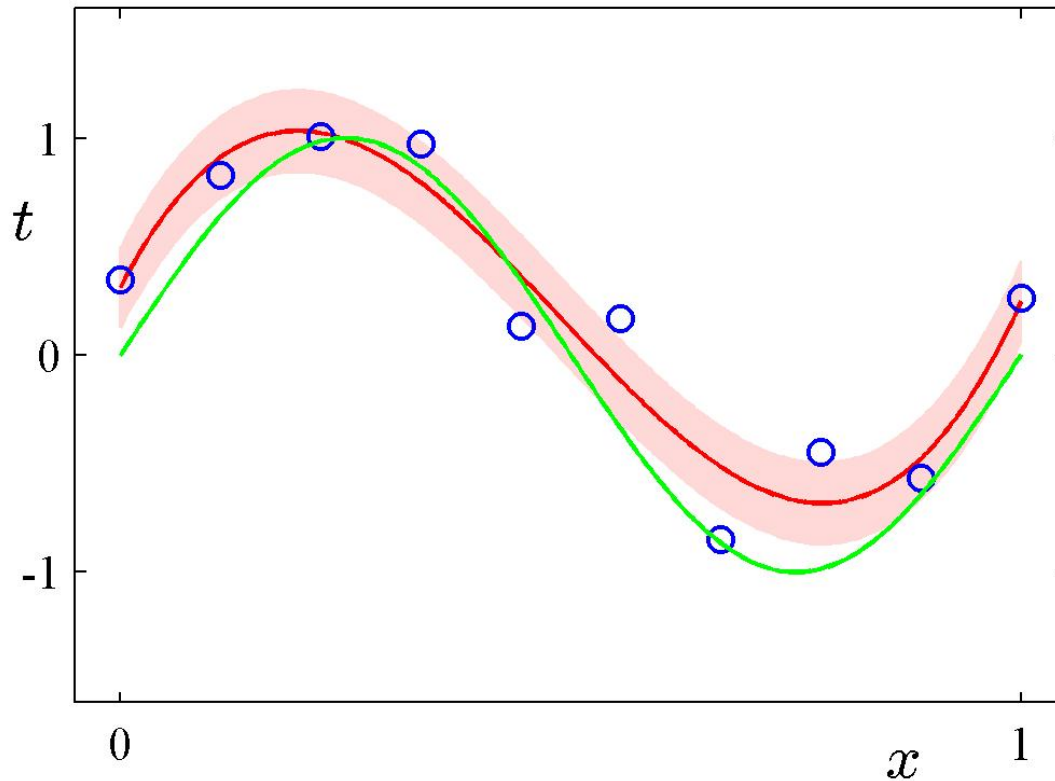
$$\mathcal{N}(t | \mathbf{m}_N^T \phi(\mathbf{x}), \sigma_N^2(\mathbf{x}))$$



$$y(x, \mathbf{w})$$

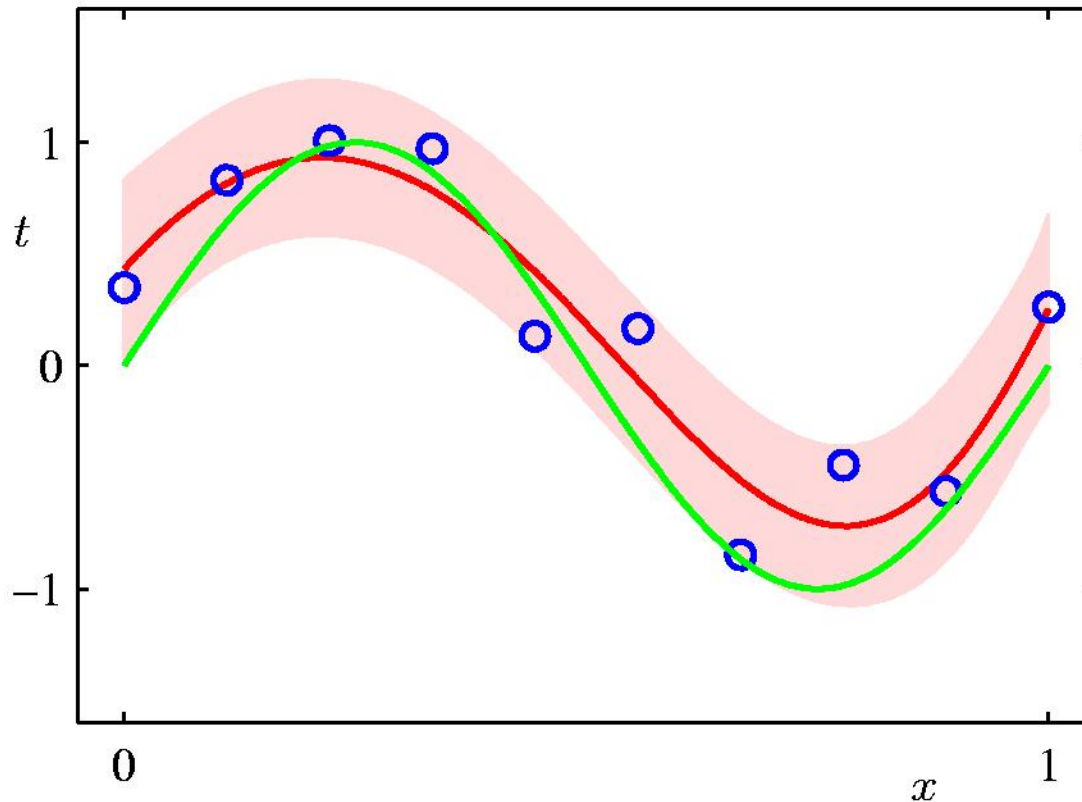
Predictive Distribution

$$p(t|x, \mathbf{w}_{\text{ML}}, \beta_{\text{ML}}) = \mathcal{N}(t|y(x, \mathbf{w}_{\text{ML}}), \beta_{\text{ML}}^{-1})$$



Bayesian Predictive Distribution

$$p(t|x, \mathbf{x}, \mathbf{t}) = \mathcal{N}(t|m(x), s^2(x))$$



The Bias-Variance Decomposition (1)

Recall the *expected squared loss*,

$$\mathbb{E}[L] = \int \{y(\mathbf{x}) - h(\mathbf{x})\}^2 p(\mathbf{x}) d\mathbf{x} + \underbrace{\iint \{h(\mathbf{x}) - t\}^2 p(\mathbf{x}, t) d\mathbf{x} dt}_{\text{noise}}$$

where

$$h(\mathbf{x}) = \mathbb{E}[t|\mathbf{x}] = \int tp(t|\mathbf{x}) dt.$$

The second term of $\mathbb{E}[L]$ corresponds to the noise inherent in the random variable t .

What about the first term?

The Bias-Variance Decomposition (2)

Suppose we were given multiple data sets, each of size N . Any particular data set, \mathcal{D} , will give a particular function $y(\mathbf{x}; \mathcal{D})$. We then have

$$\begin{aligned} & \{y(\mathbf{x}; \mathcal{D}) - h(\mathbf{x})\}^2 \\ &= \{y(\mathbf{x}; \mathcal{D}) - \mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})] + \mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})] - h(\mathbf{x})\}^2 \\ &= \{y(\mathbf{x}; \mathcal{D}) - \mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})]\}^2 + \{\mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})] - h(\mathbf{x})\}^2 \\ &\quad + 2\{y(\mathbf{x}; \mathcal{D}) - \mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})]\}\{\mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})] - h(\mathbf{x})\}. \end{aligned}$$

The Bias-Variance Decomposition (3)

Taking the expectation over \mathcal{D} yields

$$\begin{aligned} & \mathbb{E}_{\mathcal{D}} [\{y(\mathbf{x}; \mathcal{D}) - h(\mathbf{x})\}^2] \\ &= \underbrace{\{\mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})] - h(\mathbf{x})\}^2}_{(\text{bias})^2} + \underbrace{\mathbb{E}_{\mathcal{D}} [\{y(\mathbf{x}; \mathcal{D}) - \mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})]\}^2]}_{\text{variance}}. \end{aligned}$$

The Bias-Variance Decomposition (4)

Thus we can write

$$\text{expected loss} = (\text{bias})^2 + \text{variance} + \text{noise}$$

where

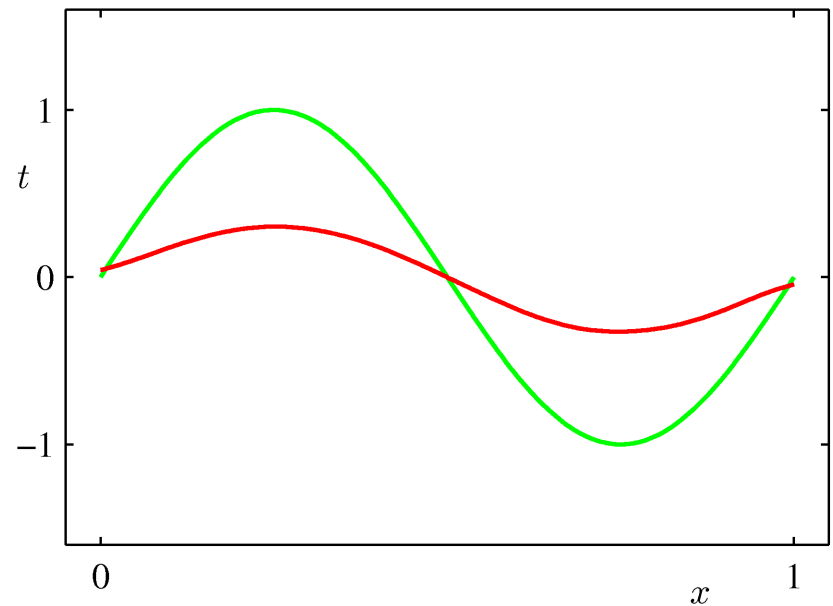
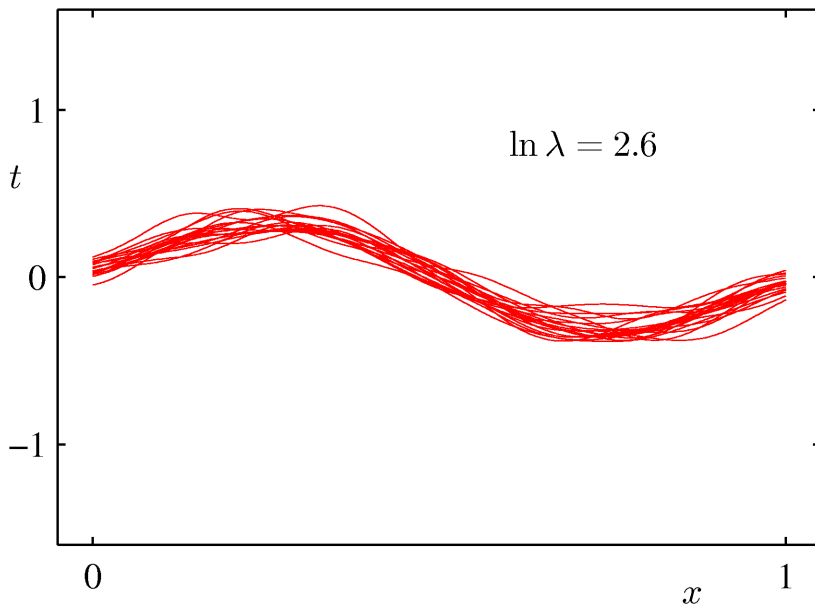
$$(\text{bias})^2 = \int \{\mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})] - h(\mathbf{x})\}^2 p(\mathbf{x}) \, d\mathbf{x}$$

$$\text{variance} = \int \mathbb{E}_{\mathcal{D}} [\{y(\mathbf{x}; \mathcal{D}) - \mathbb{E}_{\mathcal{D}}[y(\mathbf{x}; \mathcal{D})]\}^2] p(\mathbf{x}) \, d\mathbf{x}$$

$$\text{noise} = \iint \{h(\mathbf{x}) - t\}^2 p(\mathbf{x}, t) \, d\mathbf{x} \, dt$$

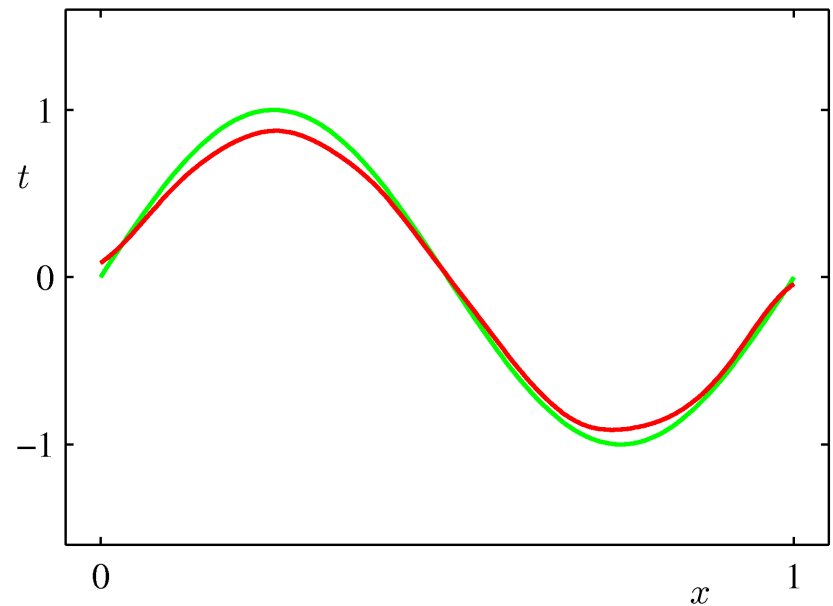
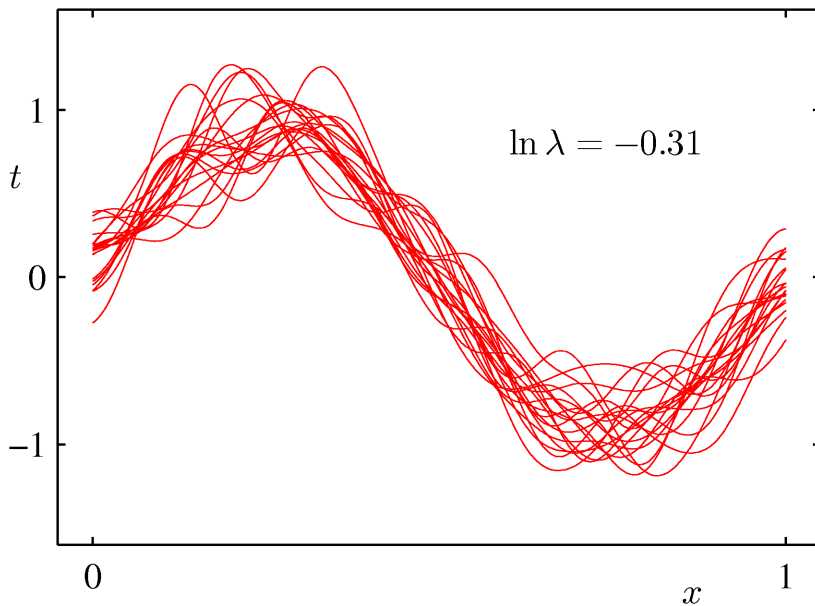
The Bias-Variance Decomposition (5)

Example: 25 data sets from the sinusoidal, varying the degree of regularization, λ .



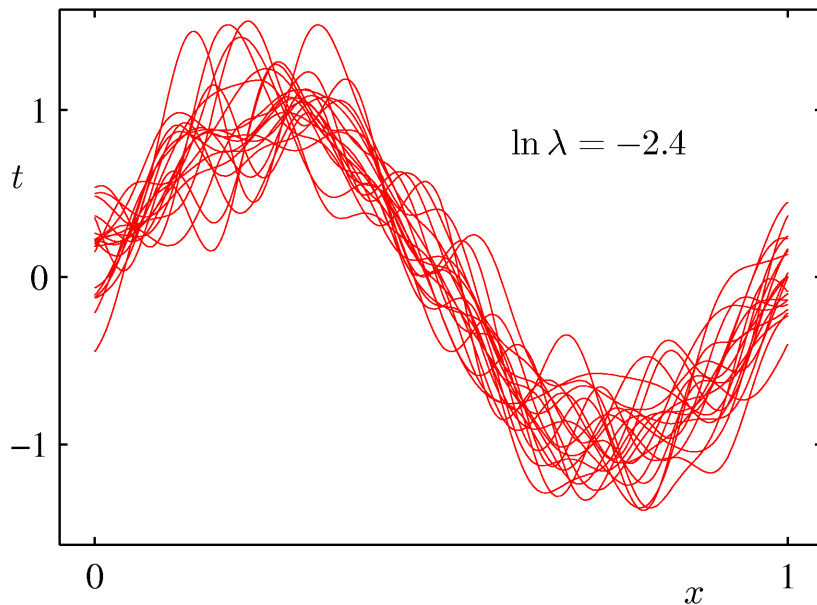
The Bias-Variance Decomposition (6)

Example: 25 data sets from the sinusoidal, varying the degree of regularization, λ .



The Bias-Variance Decomposition (7)

Example: 25 data sets from the sinusoidal, varying the degree of regularization, λ .



The Bias-Variance Trade-off

From these plots, we note that an over-regularized model (large λ) will have a high bias, while an under-regularized model (small λ) will have a high variance.

